

Session 15

Recall...

Ex. A  $Y = aX + b$ ,  $a, b \in \mathbb{R}$

15.1

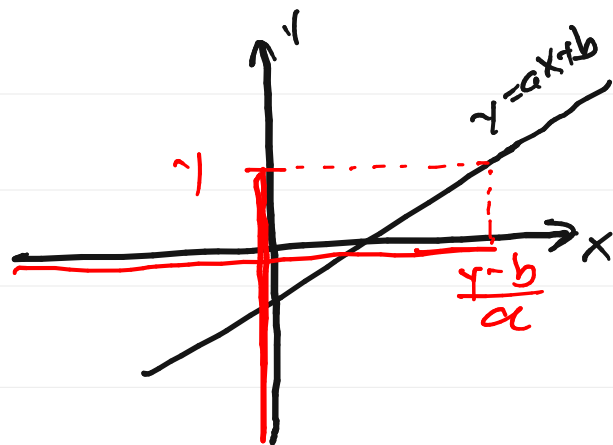
$\Rightarrow g(x) = ax + b$  (affine transformation)

two cases:  $a \gtrless 0$ .

(i)  $a > 0$

$$\begin{aligned} F_Y(y) &= P(\{ \Xi Y \leq y \}) = P(\{ \Xi aX + b \leq y \}) \\ &= P(\{ \Xi X \leq \frac{y-b}{a} \}) \\ &= F_X\left(\frac{y-b}{a}\right) \end{aligned}$$

$$\begin{aligned} f_Y(y) &= \frac{dF_Y(y)}{dy} = f_X\left(\frac{y-b}{a}\right) \cdot \frac{1}{a} \\ &= \frac{1}{a} f_X\left(\frac{y-b}{a}\right) \end{aligned}$$



Recall...

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$$(ii) \underline{a < 0}: F_{Y'}(y) = P(\sum_{i=1}^n Y_i \leq y)$$

$$= P(\sum_{i=1}^n aX_i + b \leq y)$$

$$= P(\sum_{i=1}^n X_i \geq \frac{y-b}{a})$$

$$= P(\sum_{i=1}^n X_i \geq \frac{y-b}{a})$$

$$= 1 - F_{\#} \left( \frac{y-b}{a} \right)$$

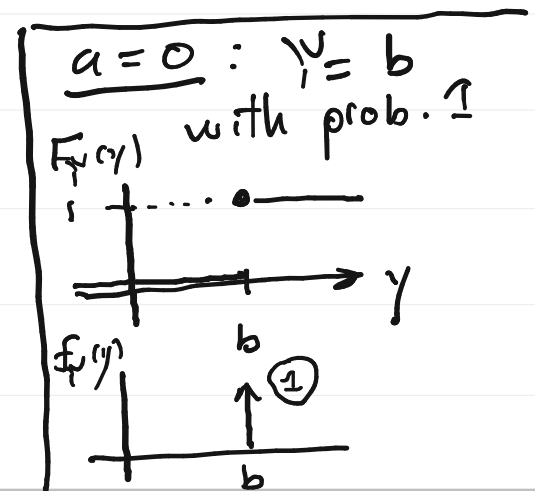
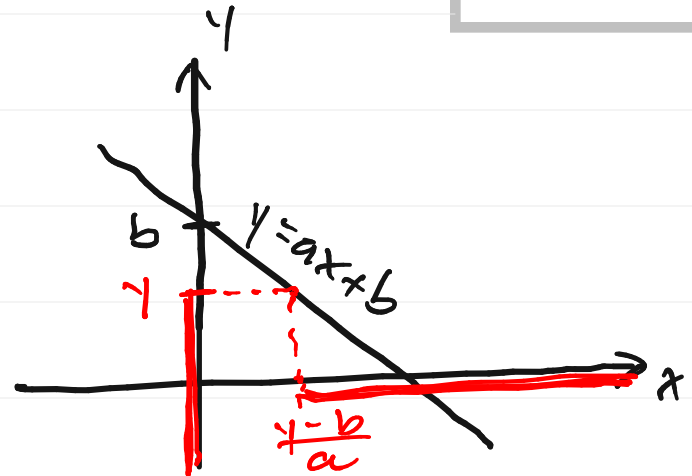
$$\Rightarrow f_{Y'}(y) = \frac{dF_{Y'}(y)}{dy} = \frac{d}{dy} \left[ 1 - F_{\#} \left( \frac{y-b}{a} \right) \right]$$

$$= -f_{\#} \left( \frac{y-b}{a} \right) \cdot \frac{1}{a}$$

$$= \frac{1}{a} f_{\#} \left( \frac{y-b}{a} \right)$$

combining (i) and (ii):

$$f_{Y'}(y) = \frac{1}{|a|} f_{\#} \left( \frac{y-b}{a} \right)$$



Ex. B  $Y = g(X) = X^2 \Rightarrow g(x) = x^2.$

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Note immediately that  $F_Y(y) = 0, y < 0.$

For  $y > 0$ :

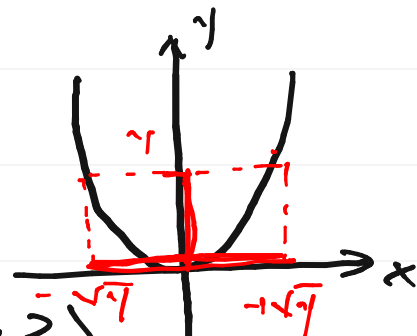
$$F_Y(y) = P(\{Y \leq y\}) = P(\{X^2 \leq y\}).$$

$$= P(\{-\sqrt{y} \leq X \leq \sqrt{y}\})$$

$$= P(\{-\sqrt{y} \leq X \leq \sqrt{y}\})$$

$$= F_X(\sqrt{y}) - F_X(-\sqrt{y})$$

$$\therefore F_Y(y) = [F_X(\sqrt{y}) - F_X(-\sqrt{y})] \cdot \mathbb{1}_{(0, \infty)}(y)$$



$$\begin{aligned}f_{\gamma}(\gamma) &= \frac{dF_{\gamma}(\gamma)}{d\gamma} \\&= f_{\#}(\sqrt{\gamma}) \cdot \frac{1}{2\sqrt{\gamma}} - f_{\#}(-\sqrt{\gamma}) \left(-\frac{1}{2\sqrt{\gamma}}\right) \\&= \frac{1}{2\sqrt{\gamma}} [f_{\#}(\sqrt{\gamma}) + f_{\#}(-\sqrt{\gamma})], \gamma > 0 \\&= \frac{1}{2\sqrt{\gamma}} [f_{\#}(\sqrt{\gamma}) + f_{\#}(-\sqrt{\gamma})] \cdot \frac{1(\gamma)}{(0, \infty)}.\end{aligned}$$

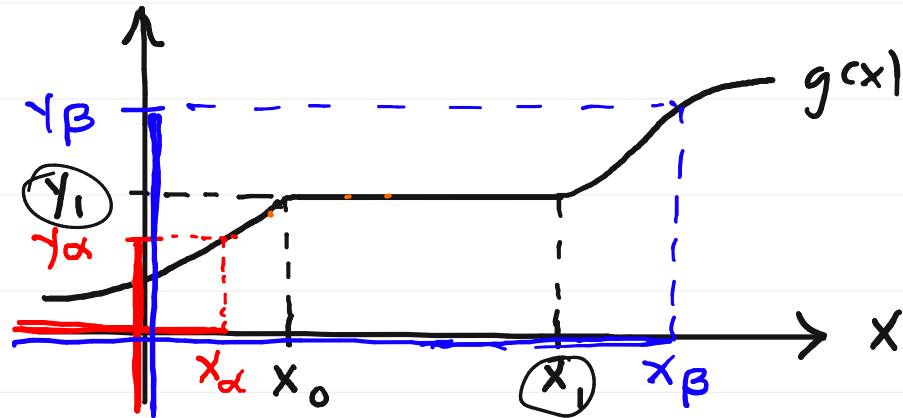
Ex. 2 Suppose that  $g(x)$  is constant across an interval  $[x_0, x_1]$

$$F_Y(y) = P(\{Y \leq y\})$$

$$(i) F_Y(y_\alpha) = P(\{Y \leq y_\alpha\}) \\ = P(\{X \leq x_\alpha\})$$

$$= P(\{X \leq g^{-1}(y_\alpha)\})$$

$$= F_X(g^{-1}(y_\alpha))$$



$$= P(\{X \in g^{-1}(-\infty, y_\alpha]\})$$

$$(ii) \underline{y > y_1}: F_Y(y_\beta) = P(\{Y \leq y_\beta\}) = P(\{X \leq x_\beta\})$$

$$= P(\{X \leq g^{-1}(y_\beta)\})$$

$$= F_X(g^{-1}(y_\beta))$$

$$F_Y(y_1) = P(\{Y \leq y_1\}) = P(\{X \leq x_1\})$$

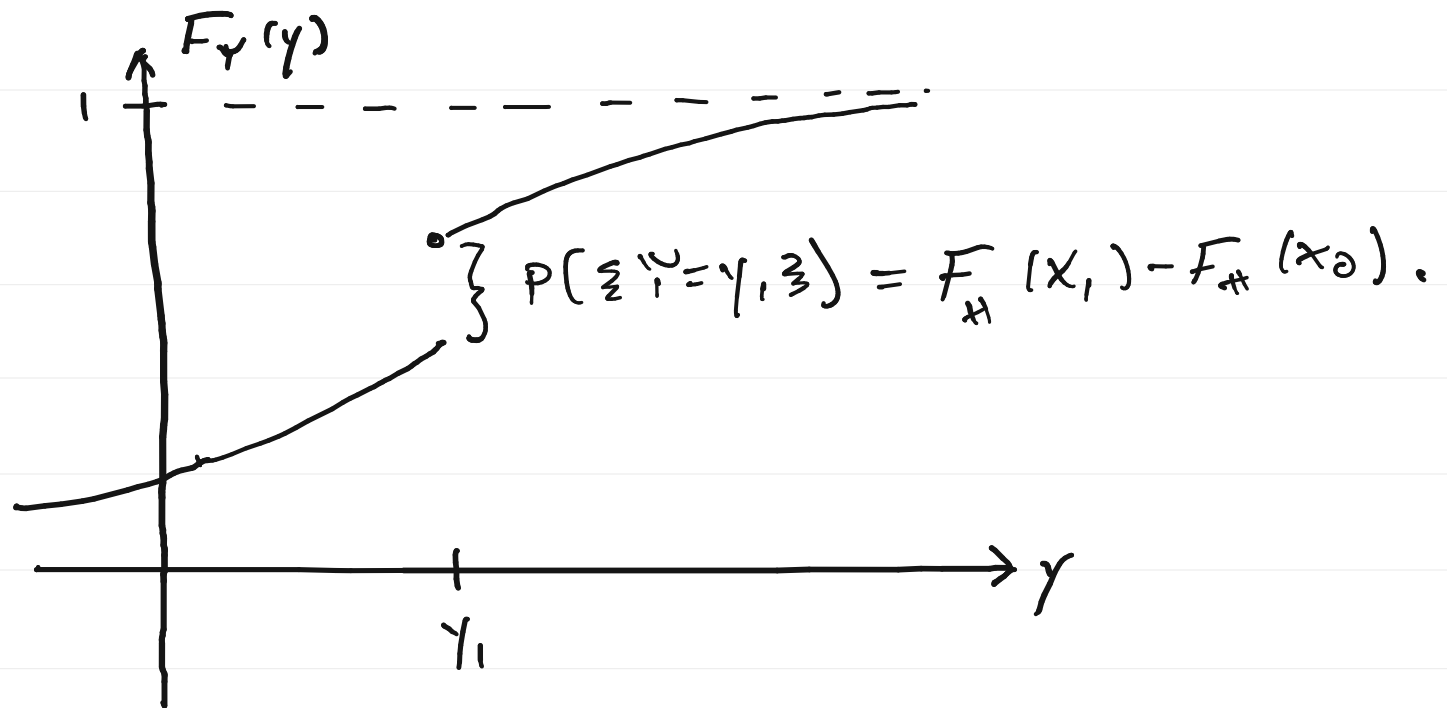
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$$= P(\{X \leq x_0\}) + P(\{x_0 < X \leq x_1\})$$

$$= F_X(x_0) + [F_X(x_1) - F_X(x_0)]$$

$$= F_X(x_1) \quad \underbrace{\hspace{10em}}_{\geq 0}$$

$$\therefore F_Y(y) = \begin{cases} F_X(g^{-1}(y)), & y < y_1 \\ F_X(x_1), & y = y_1 \\ F_X(g^{-1}(y)), & y > y_1 \end{cases}$$

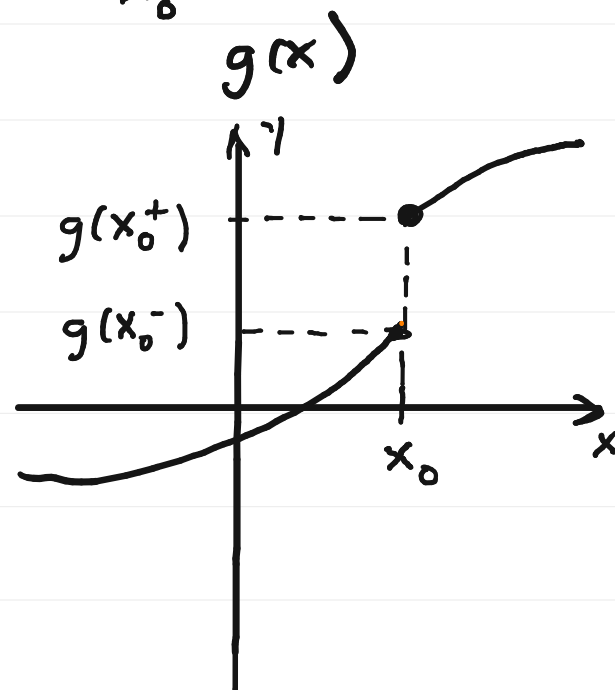




Ex. 3 Assume that  $g(x)$  has a jump discontinuity at  $x_0$

$$g(x_0^+) \neq g(x_0^-)$$

Assume  $g(x) < g(x_0^-), x < x_0$   
 $g(x) > g(x_0^+), x > x_0$



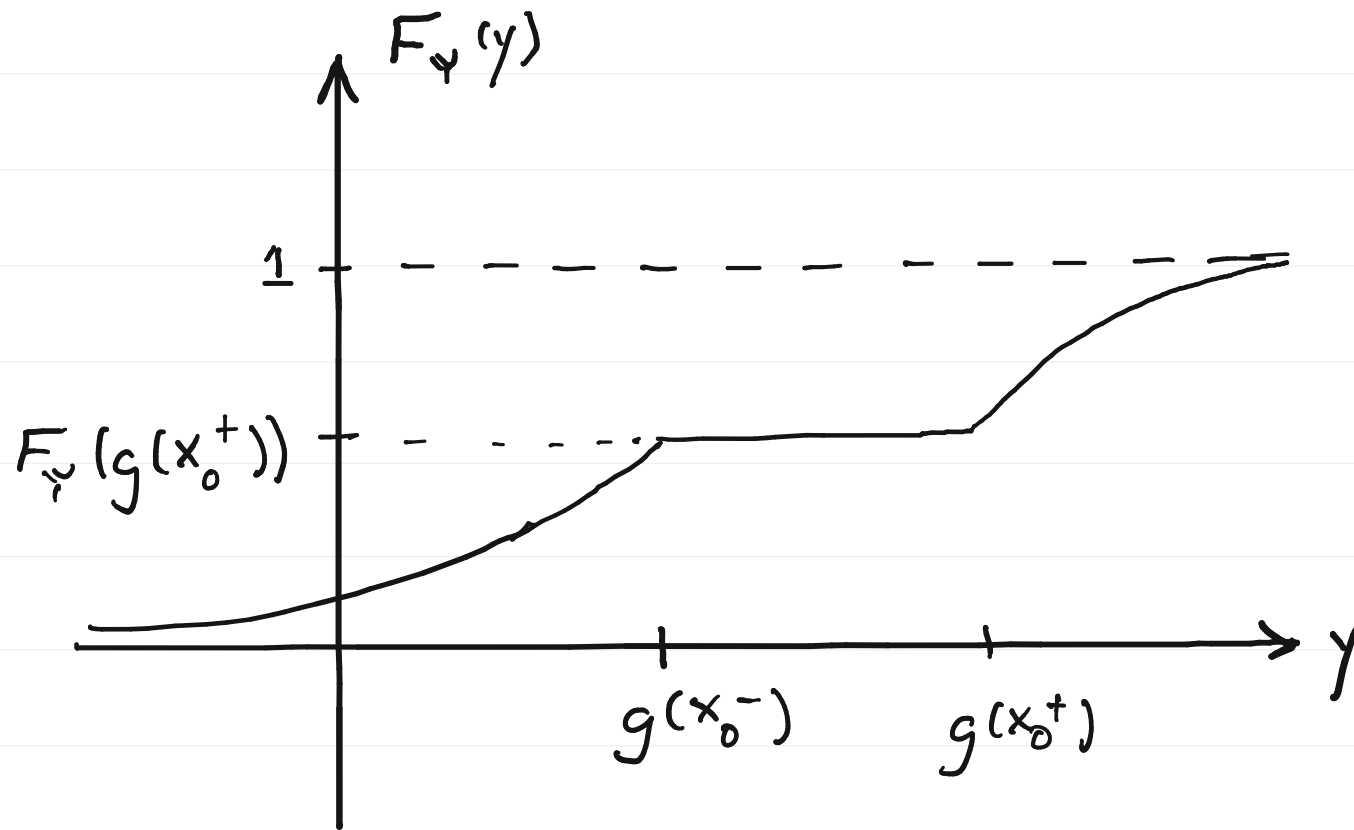
If  $y: g(x_0^-) \leq y \leq g(x_0^+)$ :

$$\begin{aligned} F_{\Psi}(y) &= P(\{ \Psi \leq y \}) = P(\{ X \leq x_0^+ \}) = P(\{ X \leq x_0^- \}) \\ &= F_X(x_0^+) = F_X(x_0^-) \end{aligned}$$

if  $X$  is absolutely continuous.

$\therefore F_{\Psi}(y)$  appears as follows:

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# The Direct pdf Method

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Suppose  $Y = g(X)$ , where  $g: \mathbb{R} \rightarrow \mathbb{R}$   
such that the inverse  $g^{-1}(\cdot)$  exists,

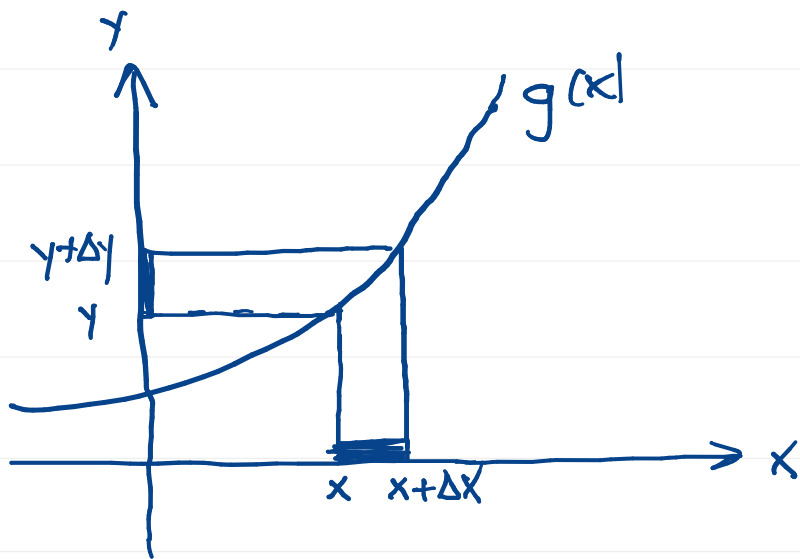
(i.e.,  $\exists y = g(x) \Rightarrow x = g^{-1}(y)$  is unique.),

and assume that  $x(y) = g^{-1}(y)$

$$\frac{dx}{dy} = \frac{dg^{-1}(y)}{dy} \text{ exists}$$

Then

$$f_Y(y) = f_X(g^{-1}(y)) \cdot \left| \frac{dg^{-1}(y)}{dy} \right|$$
$$= f_X(x(y)) \cdot \left| \frac{dx(y)}{dy} \right|, \text{ where } x(y) = g^{-1}(y).$$



$$P(\{X \in [x, x + \Delta x)\}) = P(\{Y \in [y, y + \Delta y]\})$$

$$\parallel$$

$$f_X(x) \cdot \Delta x = f_Y(y) \cdot \Delta y$$

$$\parallel$$

$$\Rightarrow f_Y(y) = f_X(x) \cdot \frac{\Delta x}{\Delta y} \approx f_X(x) \frac{dx}{dy}$$

Example: Let  $X \sim U[0,1]$  and let

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$$Y = g(X) = \sqrt{X}$$

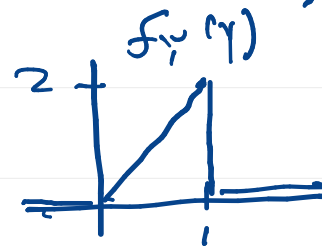
Find  $f_Y(y)$ .  $f_X(x) = 1_{[0,1]}(x)$

Solution:  $y = g(x) = \sqrt{x} \Rightarrow x = y^2$   
 $x(y) = y^2$

$$\frac{dx(y)}{dy} = \frac{dy^2}{dy} = 2y$$

$$\Rightarrow f_Y(y) = f_X(x(y)) \cdot |2y|$$

$$= 1_{[0,1]}(y^2) \cdot 2y = 2y \cdot 1_{[0,1]}(y^2)$$



Example: Let  $X$  be a Gaussian RV  
with pdf

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}.$$

Let  $Y = aX + b$ . Find  $f_Y(y)$ .

Solution:  $Y = aX + b$

$$f_Y(y) = f_X(x(y)) \left| \frac{dx(y)}{dy} \right| \quad \begin{array}{l} y = ax + b \\ x = \frac{y-b}{a} \end{array}$$

$$\frac{dx}{dy} = \frac{1}{a} \quad x(y) = \frac{y-b}{a}$$

$$f_Y(y) = \frac{1}{\sqrt{2\pi}} \exp \left\{ -\frac{(y-b)^2}{2a^2} \right\} \cdot \frac{1}{|a|}$$

$$= \frac{1}{\sqrt{2\pi}|a|} \exp \left\{ -\frac{(y-b)^2}{2a^2} \right\}.$$

# Mean, Variance and Expectation

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Defn: The mean or expected value  
of a RV  $X$  with pdf  $f_X(x)$

is

$$E[X] \triangleq \int_{-\infty}^{\infty} x f_X(x) dx.$$

What about discrete RVs.

n.b. The definition above also applies to discrete RVs if we write their pdf using  $\delta$ -functions:

If  $P(\{X = x_k\}) = P_X(x_k) = p_k$   
over a discrete index set

$$f_X(x) = \sum_k P_X(x_k) \delta(x - x_k) = \sum_k p_k \delta(x - x_k)$$

and

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_{-\infty}^{\infty} x \left( \sum_k p_k \delta(x - x_k) \right) dx$$

$$= \sum_k p_k \underbrace{\int_{-\infty}^{\infty} x \delta(x - x_k) dx}_{x_k} = \sum_k p_k x_k$$

$$= \sum_k P_X(x_k) \cdot x_k$$



∴ For a discrete RV  $X$ , we have

$$E[X] = \sum_k x_k P_X(x_k).$$

If you know about Riemann-Stieltjes integrals, you can write

$$E[X] = \int_{-\infty}^{\infty} x dF_X(x)$$

$$= \begin{cases} \sum_k x_k P_X(x_k) & (\text{discrete RV } X) \\ \int_{-\infty}^{\infty} x f_X(x) dx & (\text{continuous RV } X) \end{cases}$$

Defn: Let  $X$  be a RV on  $(\Omega, \mathcal{F}, P)$

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and let  $M \in \mathcal{F}$ . Then the conditional mean of  $X$  conditioned on  $M$  is

$$E[X|M] \triangleq \int_{-\infty}^{\infty} x f_X(x|M) dx.$$

(n.b. If  $X$  is discrete, we have the conditional

pmf  $P_X(x_k|M) = P(\{X=x_k\}|M)$ , and then

$$\begin{aligned} E[X|M] &= \int_{-\infty}^{\infty} x f_X(x|M) dx = \int_{-\infty}^{\infty} x \left( \sum_k P_X(x_k|M) \delta(x-x_k) \right) dx \\ &= \sum_k P_X(x_k|M) \cdot \int_{-\infty}^{\infty} x \delta(x-x_k) dx = \sum_k x_k P_X(x_k|M) \end{aligned}$$

Example: Let  $X$  be an exponentially distributed RV with pdf

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$$f_X(x) = \frac{1}{\mu} \exp\left\{-\frac{x}{\mu}\right\} \mathbb{1}_{[0, \infty)}(x), \mu > 0$$

What is  $E[X]$ ?

$$\begin{aligned} E[X] &= \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^{\infty} x \cdot \frac{1}{\mu} e^{-x/\mu} dx \\ &= \dots = \left[ -x e^{-x/\mu} - \mu e^{-x/\mu} \right]_0^{\infty} = \boxed{\mu} \end{aligned}$$