

## Session 23

### Iterated Expectation

23.1

Another common situation:

$$\begin{aligned} E[g(X, Y)] &= \iint_{\mathbb{R}^2} g(x, y) f_{X, Y}(x, y) dx dy \\ &= \iint_{\mathbb{R}^2} g(x, y) f_Y(y|x) f_X(x) dx dy \\ &= \int_{\mathbb{R}} f_X(x) \left[ \int_{\mathbb{R}} g(x, y) f_Y(y|x) dy \right] dx \\ &= \int_{\mathbb{R}} f_X(x) E[g(X, Y) | X=x] dx = \dots \\ & (= E[\varphi(X)].) \end{aligned}$$

$$\dots = \int_{-\infty}^{\infty} f_{\#}(x) \underbrace{E[g(\#,\psi) | \# = x]}_{\varphi(x)} dx$$

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$$= \int_{-\infty}^{\infty} f_{\#}(x) \varphi(x) dx = E_{\#}[ \varphi(\#) ] \leftarrow$$

$$\text{where } \varphi(x) = E[g(\#, \psi) | \# = x]$$

$$\therefore E[g(\#, \psi)] = E_{\#}[ \varphi(\#) ]$$

$$= E_{\#}[ E_{\psi}[g(\#, \psi) | \#] ]$$

$$= E[ E[g(\#, \psi) | \#] ]$$

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Summarizing, we have

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$$E_{\#,\psi}[g(\#, \psi)] = E_{\#}[ E_{\psi}[g(\#, \psi) | \#] ]$$

n.b. The terminology "iterated" comes from "iterated integration"

$$E[g(\#, \psi)] = \int_{-\infty}^{\infty} f_{\#}(x) \underbrace{\int_{-\infty}^{\infty} g(x, \psi) f_{\psi}(\psi | \# = x) d\psi}_{E_{\psi}[g(\#, \psi) | \# = x]} dx$$

One very important application of iterated expectation is Minimum Mean-Square estimation.

So we have

$$\begin{aligned}
 E[g(X, Y)] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{X, Y}(x, y) dx dy \\
 &= \dots = \int_{-\infty}^{\infty} f_X(x) \underbrace{\int_{-\infty}^{\infty} g(x, y) f_Y(y | \{X=x\}) dy}_{E[g(X, Y) | \{X=x\}] = \varphi(x)} dx \\
 &= E_X [E_Y [g(X, Y) | X]] = E_X [\varphi(X)].
 \end{aligned}$$

### Minimum Mean-Square Estimation

- Let  $X$  and  $Y$  be two  $j$ -dist RVs with  $j$ -pdf  $f_{X, Y}(x, y)$ .
- Suppose we want to estimate the value of  $Y$  given that we have observed  $\{X=x\}$ .

Q: What is the best estimate of the value of  $Y$  given that we know  $X=x$ ?

What do we mean by best?

One commonly used error criterion is square error.  
 $\Rightarrow$  Design the estimator that minimizes mean-square-error.

We want to find the function  $c(x)$  to estimate the value of  $Y$  given that we observe  $X=x$  such that

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$$\mathcal{E} = E[(Y - c(X))^2]$$

is minimized.

Claim: The mean-square error  $\mathcal{E}$  is minimized when

$$c(x) = E[Y | \{X=x\}].$$

Proof:  $\mathcal{E} = E[(Y - c(X))^2]$

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$$\begin{aligned} &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (y - c(x))^2 f_{X,Y}(x,y) dx dy \\ &= \int_{-\infty}^{\infty} f_X(x) \left[ \int_{-\infty}^{\infty} (y - c(x))^2 \cdot f_Y(y|x) dy \right] dx \\ &\quad \underbrace{\hspace{10em}}_{\varphi(x)} \end{aligned}$$

$\mathcal{E}$  is minimized if we pick  $c(x)$  such that the inner integral is minimized for each value of  $x$ .

We minimize the inner integral as follows (for any particular  $x$ ):

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$$\begin{aligned} & \frac{\partial}{\partial c(x)} \left\{ \int_{-\infty}^{\infty} [y - c(x)]^2 f_Y(y|x) dy \right\} \\ &= -2 \int_{-\infty}^{\infty} [y - c(x)] f_Y(y|x) dy = 0 \\ \Rightarrow & \int_{-\infty}^{\infty} (y - c(x)) f_Y(y|x) dy = 0 \\ \Rightarrow & \underbrace{\int_{-\infty}^{\infty} y f_Y(y|x) dy}_{E[Y|X=x]} - c(x) \underbrace{\int_{-\infty}^{\infty} f_Y(y|x) dy}_1 = 0 \end{aligned}$$

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$$\Rightarrow c(x) = E[Y|X=x]$$

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We will use the following notation for the MMSE estimator

$$\hat{Y}_{\text{MMS}}(x) = E[Y|X=x]$$

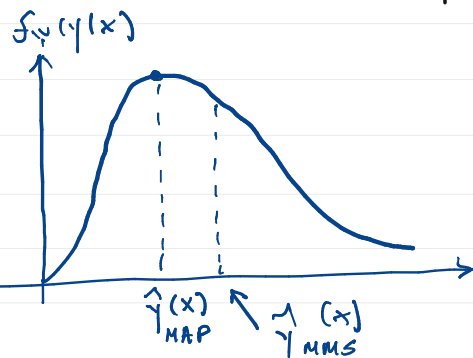
Similarly, by symmetry

$$\hat{X}_{\text{MMS}}(y) = E[X|Y=y] \quad (\text{Exercise})$$

Consider another estimator:

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$$\hat{\gamma}_{\text{MAP}}(x) = \arg \max_{\gamma} \{ f_{\gamma}(y|x) \}$$



MAP  $\triangleq$  Maximum  
Aposteriori  
Probability

The MAP estimators of interest are

$$\hat{\gamma}_{\text{MAP}}(x) = \arg \max_{\gamma} \{ f_{\gamma}(y|x=x) \}$$

$$\hat{x}_{\text{MAP}}(y) = \arg \max_{x} \{ f_{x}(x|y=y) \}$$

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Papoulis in the reading discussed  
the Linear Minimum Mean Square Error  
(LMMSE) estimator

$$c(x) = \underline{ax+b} \quad \text{assumed form.}$$

Don't worry about this. You  
are not responsible for it

## Random Vectors

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- We've considered two RVs on  $(\Omega, \mathcal{F}, P)$ .
- We can extend this to  $n$  RVs on  $(\Omega, \mathcal{F}, P)$ :

$$X_1(\omega), X_2(\omega), \dots, X_n(\omega).$$

- We can arrange these RVs as elements of a vector.

(A random vector.)

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Defn: Let  $X_1, \dots, X_n$  be  $n$  jointly distributed RVs on  $(\Omega, \mathcal{F}, P)$ .

Then the vector of RVs

$$\underline{X} = (X_1, \dots, X_n) \leftarrow \text{Row Vector}$$

is a random vector (RVec) defined on  $(\Omega, \mathcal{F}, P)$ .

Alternatively, we can think of a RVec as a mapping from  $\Omega$  to  $\mathbb{R}^n$

$$\underline{X} : \Omega \rightarrow \mathbb{R}^n$$

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$$\begin{aligned}
 \underline{\text{CDF:}} \quad F_{\underline{X}}(\underline{x}) &= F_{\underline{X}}(x_1, \dots, x_n) \\
 &= F_{X_1, \dots, X_n}(x_1, \dots, x_n) \\
 &= P(\{X_1 \leq x_1\} \cap \{X_2 \leq x_2\} \cap \dots \cap \{X_n \leq x_n\})
 \end{aligned}$$

$$\begin{aligned}
 \underline{\text{PDF:}} \quad f_{\underline{X}}(\underline{x}) &= f_{\underline{X}}(x_1, \dots, x_n) \\
 &= f_{X_1, \dots, X_n}(x_1, \dots, x_n) \\
 &= \frac{\partial^n F_{\underline{X}}(x_1, \dots, x_n)}{\partial x_1 \partial x_2 \dots \partial x_n}
 \end{aligned}$$

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Let  $D \subset \mathbb{R}^n$  ( $D \in \mathcal{B}(\mathbb{R}^n)$ )

$$\begin{aligned}
 \text{Then } P(\{\underline{X} \in D\}) &= \int_D f_{\underline{X}}(\underline{x}) d\underline{x} \\
 &= \int_{\mathbb{R}^n} f_{\underline{X}}(\underline{x}) \cdot \mathbb{1}_D(\underline{x}) d\underline{x} \\
 &= \underbrace{\int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty}}_{n\text{-fold integration}} f_{\underline{X}}(x_1, \dots, x_n) \cdot \frac{\mathbb{1}_D((x_1, \dots, x_n))}{dx_1 \dots dx_n}
 \end{aligned}$$



All the general properties of  
cdf's and pdf's generalize from  
2-dimensions to n-dimensions:

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e.g. Given  $j$ -dist RVs  $X_1, X_2, X_3, X_4$

$$F_{X_1, X_3}(x_1, x_3) = F_{X_1, X_2, X_3, X_4}(x_1, +\infty, x_3, +\infty).$$

or

$$f_{X_1, X_3}(x_1, x_3) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X_1, X_2, X_3, X_4}(x_1, x_2, x_3, x_4) dx_2 dx_4$$

### Transformations on RVecs

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Given  $\underline{X} = (X_1, \dots, X_n) \sim \text{RVec}$

we can define a new RVec

$$\underline{Y} = (Y_1, \dots, Y_k)$$

where  $Y_j = g_j(\underline{X})$ ,  $j = 1, \dots, k$

where  $k \begin{matrix} > \\ = \\ < \end{matrix} n$

How do we find  $F_{\underline{Y}}(y)$  or  $f_{\underline{Y}}(y)$ ?

Given  $\underline{X} = (X_1, \dots, X_n) \sim \text{RVec}$

we can define the new RVec

$$\underline{Y} = (Y_1, \dots, Y_k),$$

where

$$Y_j = g_j(\underline{X}), \quad j = 1, \dots, k.$$

Here  $k \geq n$ .

How do we find  $F_{\underline{Y}}(y)$  or  $f_{\underline{Y}}(y)$ ?

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To find  $F_{\underline{Y}}(y_1, \dots, y_k)$ , define

$$D(y_1, \dots, y_k) \triangleq \{(x_1, \dots, x_n) \in \mathbb{R}^n : g_1(x_1, \dots, x_n) \leq y_1, \dots, \\ g_k(x_1, \dots, x_n) \leq y_k\} \subset \mathbb{R}^n \quad (\in \mathcal{B}(\mathbb{R}^n))$$

Then

$$F_{\underline{Y}}(y_1, \dots, y_k) = \int \dots \int_{D(y_1, \dots, y_k)} f_{\underline{X}}(x_1, \dots, x_n) dx_1 \dots dx_n$$

and

$$f_{\underline{Y}}(y_1, \dots, y_k) = \frac{\partial^k F_{\underline{Y}}(y_1, \dots, y_k)}{\partial y_1 \partial y_2 \dots \partial y_k}.$$

## Direct Approach ( $k=n$ and one-to-one mapping)

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Theorem: Given RVec  $\underline{X} = (X_1, \dots, X_n)$ ,  
define a new RVec  $\underline{Y} = (Y_1, \dots, Y_n)$   
by the one-to-one mapping

$$G: \mathbb{R}^n \rightarrow \mathbb{R}^n$$

described by

$$Y_1 = g_1(\underline{X}), \dots, Y_n = g_n(\underline{X}).$$

Let  $X_1 = h_1(\underline{Y}) = x_1(\underline{Y}), \dots,$   
 $X_n = h_n(\underline{Y}) = x_n(\underline{Y}).$

Then the  $j$ -pdf of  $\underline{Y}$  is

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$$f_{\underline{Y}}(y_1, \dots, y_n) = f_{\underline{X}}(x_1(y), \dots, x_n(y)) \left| \frac{\partial(x_1, \dots, x_n)}{\partial(y_1, \dots, y_n)} \right|$$

where the Jacobian is given by

$$\frac{\partial(x_1, \dots, x_n)}{\partial(y_1, \dots, y_n)} = \begin{vmatrix} \frac{\partial x_1}{\partial y_1} & \frac{\partial x_1}{\partial y_2} & \dots & \frac{\partial x_1}{\partial y_n} \\ \frac{\partial x_2}{\partial y_1} & \frac{\partial x_2}{\partial y_2} & \dots & \frac{\partial x_2}{\partial y_n} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial x_n}{\partial y_1} & \frac{\partial x_n}{\partial y_2} & \dots & \frac{\partial x_n}{\partial y_n} \end{vmatrix}$$

n.b. If you have  $Y_1, \dots, Y_k$ , where  $k < n$ , you  
can introduce aux. variables.